

# Rao-Blackwell and sufficiency

## Recap: Rao-Blackwell

Let  $\theta$  be a parameter of interest, and  $\tau(\theta)$  some function of  $\theta$ . Let  $\hat{\tau}$  be some unbiased estimator of  $\tau(\theta)$ , and  $T$  a sufficient statistic for  $\theta$ .

Let  $\tau^* = \mathbb{E}[\hat{\tau} | T]$ . Then:

$$\textcircled{1} \quad \mathbb{E}[\tau^*] = \tau \quad (\text{unbiased})$$

$$\textcircled{2} \quad \text{Var}(\tau^*) \leq \text{Var}(\hat{\tau})$$

Reminder: If the conditional distribution of  $X_1, \dots, X_n | T$  does not depend on  $\theta$ , then  $T$  is a sufficient statistic for  $\theta$

Proof: ①  $E[\hat{\gamma}^*] = E[E[\hat{\gamma}|T]] = E[\hat{\gamma}] = \hat{\gamma}(\theta) \quad \checkmark$

(iterated expectation)

$$\begin{aligned} ② \text{Var}(\hat{\gamma}^*) &= \text{Var}(E[\hat{\gamma}|T]) \\ &\leq \text{Var}(E[\hat{\gamma}|T]) + \underbrace{E[\text{Var}(\hat{\gamma}|T)]}_{\geq 0} \\ &= \text{Var}(\hat{\gamma}) \quad (\text{law of total variance}) \end{aligned} \quad \checkmark$$

why do we need sufficiency? If  $\hat{\gamma}^*$  is a function of  $\theta$ , we can't actually calculate  $\hat{\gamma}^*$ . But if  $T$  is a sufficient statistic for  $\theta$ , then  $(X_1, \dots, X_n)|T$  does not depend on  $\theta$

$\Rightarrow \hat{\gamma}|T$  does not depend on  $\theta$

$\Rightarrow \underbrace{E[\hat{\gamma}|T]}_{\hat{\gamma}^*}$  cannot involve  $\theta$

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## Example: Rao-Blackwell

Let  $X_1, \dots, X_n \stackrel{iid}{\sim} Poisson(\lambda)$ .

a sufficient statistic:  $T = \sum_i X_i$

Consider  $\hat{\lambda} = X_1$      $E[\hat{\lambda}] = \lambda$      $\text{var}(\hat{\lambda}) = \lambda$   
 $T = \sum_i X_i$      $\lambda^* = E[\hat{\lambda} | T]$

Now,  $\underbrace{E[\sum_i X_i | T = t]}_{= \sum_i E[X_i | T = t]} = E[T | T = t] = t$

$$\Rightarrow \sum_i E[X_i | T = t] = t$$

$$\Rightarrow E[X_1 | T = t] = \frac{t}{n}$$

$$\Rightarrow \lambda^* = \frac{T}{n} = \frac{1}{n} \sum_{i=1}^n X_i$$

Conditioning on an insufficient statistic:

$X_1, \dots, X_n \stackrel{iid}{\sim} \text{Poisson}(\lambda)$

$$\hat{\lambda} = X_1$$

$\swarrow$   $X_2$  is not sufficient for  $\lambda$

$$\lambda^* = \mathbb{E}[X_1 | X_2]$$

$$= \mathbb{E}[X_1] = \lambda$$

$\uparrow$

not a valid estimator of  $\lambda$

$$x_1, x_2 \quad P(X_1=1, X_2=3, X_1+X_2=4)$$

$$= P(X_1=1, X_2=3)$$

## Example: sufficiency

Let  $X_1, \dots, X_n \stackrel{iid}{\sim} \text{Bernoulli}(p)$ .

Find a sufficient statistic for  $p$ .

$$\bar{T} = \frac{1}{n} \sum_i x_i \quad \text{or} \quad \bar{T} = \sum_i x_i \sim \text{Binomial}(n, p)$$

e.g.  $\bar{T} = \sum_i x_i$  wTS:  $f(x_1, \dots, x_n | \sum_i x_i, p)$  does not depend on  $p$

$$\begin{aligned}
 f(x_1, \dots, x_n | \sum_i x_i, p) &= \frac{f(x_1, \dots, x_n, \sum_i x_i | p)}{f(\sum_i x_i | p)} \\
 &= \frac{f(x_1, \dots, x_n | p)}{f(\sum_i x_i | p)} = \frac{\prod_i p^{x_i} (1-p)^{1-x_i}}{(\sum_i x_i) p^{\sum_i x_i} (1-p)^{n-\sum_i x_i}} \\
 &= \frac{1}{(\sum_i x_i)} \quad (\text{does not depend on } p)
 \end{aligned}$$

## Factorization theorem

Let  $x_1, \dots, x_n$  be a sample with joint probability function  $f(x_1, \dots, x_n | \theta)$ . A statistic  $T = T(x_1, \dots, x_n)$  is sufficient for  $\theta$  if and only if there exist functions  $g(t | \theta)$  and  $h(x_1, \dots, x_n)$  such that, for all possible  $x_1, \dots, x_n$  and all possible  $\theta$ ,

$$f(x_1, \dots, x_n | \theta) = \underbrace{g(T(x_1, \dots, x_n) | \theta)}_{\text{joint distribution only depends on } \theta} h(x_1, \dots, x_n)$$

## Example

Suppose  $x_1, \dots, x_n \stackrel{iid}{\sim} \text{Bernoulli}(p)$

$$\begin{aligned} f(x_1, \dots, x_n | p) &= P^{\sum_i x_i} (1-p)^{n - \sum_i x_i} \\ &= \underbrace{P^{\sum_i x_i}}_{g(\sum_i x_i | p)} \underbrace{(1-p)^{n - \sum_i x_i}}_{h(x_1, \dots, x_n)} = 1 \end{aligned}$$