

Method of moments estimators

Course so far

- + Maximum likelihood estimation
- + Logistic regression
- + Asymptotics
- + Asymptotic properties of MLEs
- + Hypothesis testing
- + Confidence intervals

Common theme: Likelihoods and MLEs

Question: Why maximum likelihood estimation?

Example

Suppose $X_1, \dots, X_n \stackrel{iid}{\sim} Uniform[0, \theta]$. How could I estimate θ ?

① $\hat{\theta}_{MLE} = X_{(n)}$

② median $U[0, \theta] = \frac{\theta}{2}$

$\hat{\theta} = \text{Sample median} \times 2$

③ $E[X] = \frac{\theta}{2} \Rightarrow \hat{\theta} = 2\bar{X}$

④ $E[X^2] = \frac{\theta^2}{12} + \frac{\theta^2}{4} = \frac{\theta^2}{3} \Rightarrow \hat{\theta} = \sqrt{\frac{3}{n} \sum_i x_i^2}$

⑤ $\hat{\theta} = 5$ (probably a terrible estimate)

$$\Rightarrow \hat{a} = \hat{\mu}_1 - \sqrt{3(\hat{\mu}_2 - \hat{\mu}_1^2)}$$

$$\hat{b} = \hat{\mu}_1 + \sqrt{3(\hat{\mu}_2 - \hat{\mu}_1^2)}$$

Example

Suppose $X_1, \dots, X_n \stackrel{iid}{\sim} \text{Uniform}[a, b]$. How could I estimate a and b ?

- ① Fix a , estimate b , alternate
- ② MLE: $\hat{a}_{MLE} = X_{(1)}$ $\hat{b}_{MLE} = X_{(n)}$

$$\text{③ } E[X] = \frac{a+b}{2} = \mu_1 \quad \hat{\mu} = \frac{1}{n} \sum_i X_i = \bar{X}$$

$$E[X^2] = \frac{1}{3}(a^2 + ab + b^2) = \mu_2 \quad \hat{\mu}_2 = \frac{1}{n} \sum_i X_i^2$$

$$b = 2\mu_1 - a$$

$$\Rightarrow \mu_2 = \frac{1}{3}(a^2 + a(2\mu_1 - a) + (2\mu_1 - a)^2)$$

$$= \frac{1}{3}(a^2 - 2a\mu_1 + 4\mu_1^2)$$

$$\Rightarrow 3\mu_2 - 3\mu_1^2 = (a - \mu_1)^2$$

$$\Rightarrow a = \mu_1 - \sqrt{3(\mu_2 - \mu_1^2)} \quad b = \mu_1 + \sqrt{3(\mu_2 - \mu_1^2)}$$

Method of moments

Let X_1, \dots, X_n be a sample from a distribution with probability function $f(x|\theta_1, \dots, \theta_k)$, with k parameters $\theta_1, \dots, \theta_k$.

$$\begin{aligned} \text{Let } \mu_1 &= E[X] = g_1(\theta_1, \dots, \theta_k) & \hat{\mu}_1 &= \frac{1}{n} \sum_i x_i \\ \mu_2 &= E[X^2] = g_2(\theta_1, \dots, \theta_k) & \hat{\mu}_2 &= \frac{1}{n} \sum_i x_i^2 \\ &\vdots & &\vdots \\ \mu_k &= E[X^k] = g_k(\theta_1, \dots, \theta_k) & \hat{\mu}_k &= \frac{1}{n} \sum_i x_i^k \end{aligned}$$

The method of moments (Mom) approach estimates $\theta_1, \dots, \theta_k$ by the solutions to

$$\hat{\mu}_1 = g_1(\hat{\theta}_1, \dots, \hat{\theta}_k)$$

$$\hat{\mu}_2 = g_2(\hat{\theta}_1, \dots, \hat{\theta}_k)$$

$$\vdots$$
$$\hat{\mu}_k = g_k(\hat{\theta}_1, \dots, \hat{\theta}_k)$$

Example

Suppose $X_1, \dots, X_n \stackrel{iid}{\sim} N(\mu, \sigma^2)$.

Find the method of moments estimates $\hat{\mu}$ and $\hat{\sigma}^2$.

$$\mu_1 = \mu$$

$$\begin{aligned}\mu_2 &= \text{Var}(X) + (\mathbb{E}[X])^2 \\ &= \sigma^2 + \mu^2\end{aligned}$$

$$\hat{\mu}_1 = \bar{X}$$

$$\hat{\mu}_2 = \frac{1}{n} \sum_i x_i^2$$

$$\sigma^2 = \mu_2 - \hat{\mu}_1^2$$

$$\Rightarrow \hat{\sigma}^2 = \hat{\mu}_2 - \hat{\mu}_1^2$$

$$= \frac{1}{n} \sum_i x_i^2 - (\bar{X})^2$$

$$= \frac{1}{n} \sum_i (x_i - \bar{X})^2$$

Example

Suppose $X_1, \dots, X_n \stackrel{iid}{\sim} Gamma(\alpha, \beta)$, i.e.

$$f(x|\alpha, \beta) = \frac{\beta^\alpha}{\Gamma(\alpha)} x^{\alpha-1} e^{-\beta x}. \text{ Then}$$

$$\mu_1 = \mathbb{E}[X] = \frac{\alpha}{\beta} \quad \mu_2 = \mathbb{E}[X^2] = \left(\frac{\alpha}{\beta}\right)^2 + \frac{\alpha}{\beta^2}$$

Use the method of moments to estimate α and β .

$$m_1 = \frac{\alpha}{\beta} \quad m_2 = \left(\frac{\alpha}{\beta} \right)^2 + \frac{\alpha}{\beta^2}$$

$$\alpha = \beta m_1$$

$$\Rightarrow m_2 = \left(\frac{\beta m_1}{\beta} \right)^2 + \frac{\beta m_1}{\beta^2}$$

$$= m_1^2 + \frac{m_1}{\beta}$$

$$\beta m_2 = \beta m_1^2 + m_1$$

$$\beta(m_2 - m_1^2) = m_1$$

$$\beta = \frac{m_1}{m_2 - m_1^2}$$

$$\Rightarrow \alpha = \frac{m_1^2}{m_2 - m_1^2}$$

$$\hat{\alpha} = \frac{\hat{m}_1^2}{\hat{m}_2 - \hat{m}_1^2}$$

$$\hat{\beta} = \frac{\hat{m}_1}{\hat{m}_2 - \hat{m}_1^2}$$

$$\begin{aligned} \hat{\alpha}, \hat{\beta} &> 0 \\ (\hat{m}_2 &> \hat{m}_1^2) \end{aligned}$$