

# Interval estimation

# Motivation

Suppose we have data  $(X_1, Y_1), \dots, (X_n, Y_n)$  with

$$Y_i \sim \text{Bernoulli}(p_i)$$

$$\log\left(\frac{p_i}{1 - p_i}\right) = \beta^T X_i$$

So far, we have discussed:

- + Finding point estimates  $\hat{\beta}$
- + Testing hypotheses about the true (but unknown) parameters  $\beta$

What are the limitations of point estimates and hypothesis tests for inference about  $\beta$ ?

# Confidence interval

```
...  
##              Estimate Std. Error z value Pr(>|z|)  
## (Intercept)  2.6415063  0.1213233   21.77  <2e-16 ***  
## WBC          -0.2892904  0.0134349  -21.53  <2e-16 ***  
## PLT          -0.0065615  0.0005932  -11.06  <2e-16 ***  
## ---  
...
```

How would I calculate a 95% confidence interval for  $\beta_1$  (the change in the log odds of dengue for a one-unit increase in WBC, holding PLT fixed)?

$$\hat{\beta}_1 \pm z_{\frac{\alpha}{2}} SE(\hat{\beta}_1) \quad \leftarrow (1-\alpha) \text{ Wald CI}$$

$$\begin{aligned} 95\% \text{ CI} &: -0.289 \pm 1.96 (0.0134) \\ &= (-0.315, -0.262) \end{aligned}$$

$\beta_1$  is either in  $(-0.315, -0.262)$  or not  
(no "95% probability" for a specific interval)

## Confidence interval

...

##	Estimate	Std. Error	z value	Pr(> z )	
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## ---					

... coverage probability  
↙

95% confidence interval for  $\beta_1$ : (-0.315, -0.262)

How do I interpret this confidence interval?

95% confident: if we take many samples and we calculate many intervals, 95% should contain the true (unknown) parameter

"we are 95% confident that a one-unit increase in WBC is associated w/ a decrease in log odds of dengue by between 0.262 and 0.315" 4/9

$$\hat{\theta} \sim N(\theta, \text{var}(\hat{\theta})) \Rightarrow \frac{\hat{\theta} - \theta}{\text{SE}(\hat{\theta})} \sim N(0, 1)$$

## Deriving the coverage probability

$(1-\alpha)$  Wald interval:  $\hat{\theta} \pm z_{\frac{\alpha}{2}} \text{SE}(\hat{\theta})$

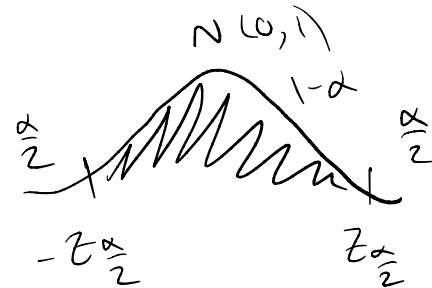
$$\hookrightarrow P(\hat{\theta} - z_{\frac{\alpha}{2}} \text{SE}(\hat{\theta}) \leq \theta \leq \hat{\theta} + z_{\frac{\alpha}{2}} \text{SE}(\hat{\theta})) = 1 - \alpha$$

endpoints are random (function of data)

$$P(\hat{\theta} - z_{\frac{\alpha}{2}} \text{SE}(\hat{\theta}) \leq \theta \leq \hat{\theta} + z_{\frac{\alpha}{2}} \text{SE}(\hat{\theta}))$$

$$= P(-z_{\frac{\alpha}{2}} \text{SE}(\hat{\theta}) \leq \hat{\theta} - \theta \leq z_{\frac{\alpha}{2}} \text{SE}(\hat{\theta}))$$

$$= P(-z_{\frac{\alpha}{2}} \leq \underbrace{\frac{\hat{\theta} - \theta}{\text{SE}(\hat{\theta})}}_{\sim N(0,1)} \leq z_{\frac{\alpha}{2}}) = 1 - \alpha$$



Note:  $\theta_0 \in [\hat{\theta} - z_{\frac{\alpha}{2}} \text{SE}(\hat{\theta}), \hat{\theta} + z_{\frac{\alpha}{2}} \text{SE}(\hat{\theta})]$

$$\Leftrightarrow \left| \frac{\hat{\theta} - \theta_0}{\text{SE}(\hat{\theta})} \right| \leq z_{\frac{\alpha}{2}}$$

i.e. the  $\alpha$ -level Wald test of  $H_0: \theta = \theta_0$  vs.  $H_A: \theta \neq \theta_0$  fails to reject 5/9

## Formal definition

Let  $\theta \in \Theta$  be a parameter of interest, and  $X_1, \dots, X_n$  a sample. Let  $C(X_1, \dots, X_n) \subseteq \Theta$  be a set constructed from  $X_1, \dots, X_n$  ( $\Rightarrow C(X_1, \dots, X_n)$  is a random set)

$C(X_1, \dots, X_n)$  is a  $1 - \alpha$  confidence set for  $\theta$  if

$$\inf_{\theta \in \Theta} P_{\theta}(\theta \in C(X_1, \dots, X_n)) = 1 - \alpha$$

(i.e.  $\forall \theta \in \Theta, P_{\theta}(\theta \in C(X_1, \dots, X_n)) \geq 1 - \alpha$ )

$\alpha$ -level test of  $H_0: \theta = \theta_0$  :  $P_{\theta_0}((X_1, \dots, X_n) \in R(\theta_0)) \leq \alpha$

## Inverting a test

↑  
rejection  
region

Theorem: Let  $\theta \in \Theta$  be a parameter of interest.

For each value of  $\theta_0 \in \Theta$ , consider testing  $H_0: \theta = \theta_0$  vs.  $H_A: \theta \neq \theta_0$ , and let  $R(\theta_0)$  be the rejection region for a level  $\alpha$  test.

Let  $C(X_1, \dots, X_n) = \{ \theta_0 : (X_1, \dots, X_n) \notin R(\theta_0) \}$

Then  $C(X_1, \dots, X_n)$  is a  $1 - \alpha$  confidence set for  $\theta$ .

Pf:  $\theta \in C(X_1, \dots, X_n) \Leftrightarrow (X_1, \dots, X_n) \notin R(\theta)$

$$\begin{aligned} \Rightarrow P_{\theta}(\theta \in C(X_1, \dots, X_n)) &= P_{\theta}((X_1, \dots, X_n) \notin R(\theta)) \\ &= 1 - P_{\theta}((X_1, \dots, X_n) \in R(\theta)) \\ &\geq 1 - \alpha \quad (\alpha\text{-level test}) // \end{aligned}$$

## Example

Suppose  $X_1, \dots, X_n \stackrel{iid}{\sim} \text{Uniform}[0, \theta]$ . We want to test

$$H_0 : \theta = \theta_0 \quad H_A : \theta \neq \theta_0$$

Find the LRT statistic for this test.

Reject  $H_0$  when  $\frac{\sup_{\theta} L(\theta | X)}{\theta} > K$

$$L(\theta | X) = \left(\frac{1}{\theta}\right)^n \mathbb{1}\{X_{(n)} \leq \theta\} \quad \hat{\theta}_{MLE} = X_{(n)}$$

$$\Rightarrow \text{reject } H_0 \text{ when } \frac{\theta_0^n}{X_{(n)}^n \mathbb{1}\{\theta_0 \geq X_{(n)}\}} > K$$

$$\Rightarrow \text{reject } H_0 \text{ when } \theta_0 < X_{(n)} \text{ or when } \frac{\theta_0}{X_{(n)}} > K^{\frac{1}{n}}$$



reject  $H_0$  when  $\theta_0 < X_{(n)}$  or  $\frac{\theta_0}{X_{(n)}} > k^{\frac{1}{n}}$

$\Rightarrow$  fail to reject  $H_0$  when  $X_{(n)} \leq \theta_0 \leq X_{(n)} k^{\frac{1}{n}}$

confidence set =  $[X_{(n)}, X_{(n)} k^{\frac{1}{n}}] = [X_{(n)}, X_{(n)} k']$

so need  $k'$  st  $P_{\theta}(\theta \in [X_{(n)}, X_{(n)} k']) \geq 1 - \alpha$

$$P_{\theta}(\theta \in [X_{(n)}, X_{(n)} k']) = P_{\theta}(X_{(n)} k' \geq \theta)$$

$$= 1 - P_{\theta}(X_{(n)} k' < \theta)$$

$$= 1 - P_{\theta}(X_{(n)} < \frac{\theta}{k'})$$

$$= 1 - \left(\frac{\theta/k'}{\theta}\right)^n = 1 - \left(\frac{1}{k'}\right)^n$$

$\Rightarrow$  choose  $k'$  st  $\left(\frac{1}{k'}\right)^n = \alpha$   $\Rightarrow k' = \frac{1}{\alpha^{\frac{1}{n}}}$   
 $1 - \alpha$  CI for  $\theta: [X_{(n)}, \frac{X_{(n)}}{\alpha^{\frac{1}{n}}}]$