

Hypothesis testing framework

$$H_0: \begin{bmatrix} \beta_3 \\ \beta_4 \end{bmatrix} = \begin{bmatrix} 0 \\ 0 \end{bmatrix}$$

← χ^2 length of part of β that we're testing

Last time

We have the model

$$Y_i \sim \text{Bernoulli}(p_i)$$

$$\log\left(\frac{p_i}{1-p_i}\right) = \beta_0 + \beta_1 \text{Sex}_i + \beta_2 \text{Age}_i + \beta_3 \text{SecondClass}_i + \beta_4 \text{ThirdClass}_i$$

We want to test whether there is a difference in the chance of survival for second and third class passengers, holding age and sex fixed.

What hypotheses should we test?

$$H_0: \beta_3 = \beta_4$$

$$\beta_4 - \beta_3 = 0$$

$$H_A: \beta_3 \neq \beta_4$$

$$\beta_4 - \beta_3 \neq 0$$

$$\text{Want: } C \text{ st } C\beta = \beta_4 - \beta_3$$

Contrasts

$$\beta_4 - \beta_3 = \underbrace{[0 \ 0 \ 0 \ -1 \ 1]}_{a^T} \begin{bmatrix} \beta_0 \\ \beta_1 \\ \vdots \\ \beta_4 \end{bmatrix} = a^T \beta$$

$$= C \beta$$

$$C = a^T$$

Def: Let $\theta \in \mathbb{R}^q$ and $a \in \mathbb{R}^q$ such that $\sum_i a_i = 0$
 Then $a^T \theta$ is called contrast

$$H_0: a^T \theta = \gamma_0$$

$$H_A: a^T \theta \neq \gamma_0$$

$$\text{ex: } H_0: a^T \beta = 0$$

$$H_A: a^T \beta \neq 0$$

$$\text{Test: } Z_n = (a^T \mathcal{I}^{-1}(\beta) a)^{-\frac{1}{2}} (a^T \hat{\beta} - 0) = \frac{a^T \hat{\beta}}{\sqrt{a^T \mathcal{I}^{-1}(\beta) a}}$$

$$\approx N(0, 1) \quad \text{under } H_0$$

$$\text{reject } H_0 \text{ when } |Z_n| > Z_{\frac{\alpha}{2}} \quad (\text{or } Z_n^2 > \chi_{1, \alpha}^2)$$

Class activity

Work on Part II from last class:

https://sta711-s23.github.io/class_activities/ca_lecture_18.html

Class activity

```
a <- c(0, 0, 0, -1, 1)
test_stat <- (t(a) %*% coef(m1))/sqrt(t(a) %*% vcov(m1) %*% a)
test_stat
```

$\hat{\beta}$

$a^T \hat{\beta}$

$\sqrt{a^T \mathcal{I}^{-1}(\beta) a}$

$\mathcal{I}^{-1}(\beta)$

```
##           [,1]
## [1,] -5.207289
```

$$Z_n = -5.21$$

$$\alpha = 0.05 \quad \Rightarrow \text{reject} \quad \text{when} \quad |Z_n| > z_{\frac{\alpha}{2}} = \underbrace{z_{0.025}}_{1.96}$$

Class activity

```
a <- c(0, 0, 0, -1, 1)
test_stat <- (t(a) %*% coef(m1))/sqrt(t(a) %*% vcov(m1) %*% a)
test_stat
```

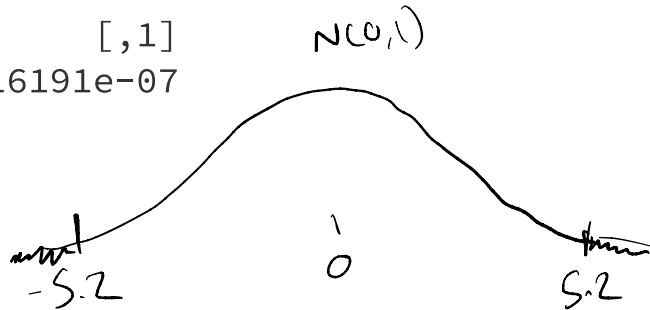
```
##           [,1]
## [1,] -5.207289
```

```
# rejection region for alpha = 0.05
qnorm(0.025, lower.tail=F)
```

```
## [1] 1.959964
```

```
# p-value
2*pnorm(abs(test_stat), lower.tail=F)
```

```
##           [,1]
## [1,] 1.916191e-07
```



Logistic regression MLE: $\text{Var}(\hat{\beta}) = V_n = \mathcal{I}^{-1}(\beta)$

Summary of Wald tests

Let $\theta \in \mathbb{R}^p$ be some parameter of interest. We wish to test the hypotheses

$$H_0 : C\theta = \gamma_0 \quad H_A : C\theta \neq \gamma_0 \quad \text{specify hypothesis}$$

for some $C \in \mathbb{R}^{q \times p}$. Given an estimator $\hat{\theta}_n$ such that

$$V_n^{-\frac{1}{2}}(\hat{\theta}_n - \theta) \xrightarrow{d} N(0, I),$$

the **Wald test** rejects when

$$\underbrace{(C\hat{\theta}_n - \gamma_0)^T (CV_n C^T)^{-1} (C\hat{\theta}_n - \gamma_0)}_{\text{calculate a test statistic}} > \underbrace{\chi_{q,\alpha}^2}_{\text{determine when to reject } H_0}$$

calculate a test
statistic

determine when to
reject H_0

General framework for hypothesis tests

Definition: Let $\theta \in \mathbb{R}^p$ be some parameter of interest. A **hypothesis** is a statement about θ . The goal of a hypothesis test is to compare two competing hypotheses:

$$H_0 : \theta \in \Theta_0 \quad H_A : \theta \in \Theta_1$$

• here $\Theta_0 \cap \Theta_1 = \emptyset$

• If $\Theta_0 = \{\theta_0\}$, H_0 is a simple hypothesis

otherwise, H_0 is a composite hypothesis (likewise for H_A)

Example :

$$H_0 : \begin{bmatrix} \beta_3 \\ \beta_4 \end{bmatrix} = \begin{bmatrix} 0 \\ 0 \end{bmatrix} \quad H_A : \begin{bmatrix} \beta_3 \\ \beta_4 \end{bmatrix} \neq 0$$

\uparrow simple null \uparrow composite alternative

Outcomes

$$H_0 : \theta \in \Theta_0 \quad H_A : \theta \in \Theta_1$$

The outcome of the test is a decision to either **reject** H_0 or **fail to reject** H_0 .

Four possibilities

Truth

H_0 is true

H_A is true

Decision

Fail to reject

Reject

Yay!	Type II error (false negative)
Type I error (false positive)	Yay!

Goal:

- 1) Control type I error
- 2) Minimize type II error

Constructing a test

$$H_0 : \theta \in \Theta_0 \quad H_A : \theta \in \Theta_1$$

observe data X_1, \dots, X_n

- ① Calculate a test statistic $T_n = T(X_1, \dots, X_n)$
- ② Choose a rejection region $R = \{(x_1, \dots, x_n) : \text{reject } H_0\}$
- ③ Reject H_0 if $(X_1, \dots, X_n) \in R$

Example : X_1, \dots, X_n iid from population w/ mean μ , variance σ^2 . $H_0: \mu = \mu_0$ $H_A: \mu \neq \mu_0$

$$\textcircled{1} \quad T(X_1, \dots, X_n) = \frac{\bar{X}_n - \mu_0}{\sigma/\sqrt{n}} = Z_n$$

$$\textcircled{2} \quad R = \left\{ (x_1, \dots, x_n) : \left| \frac{\bar{x}_n - \mu_0}{\sigma/\sqrt{n}} \right| > z_{\frac{\alpha}{2}} \right\}$$

$$\textcircled{3} \quad \text{Reject } H_0 \text{ if } (X_1, \dots, X_n) \in R, \text{ i.e. if } |Z_n| > z_{\frac{\alpha}{2}}$$

Prob. of a type I error = $P(\text{reject } H_0 \mid H_0 \text{ is true})$
Under H_0 , $Z_n \sim N(0,1) \Rightarrow \text{prob. of type I error} = P(|Z| > z_{\frac{\alpha}{2}}) = \alpha$

Power function

Suppose we reject H_0 when $(X_1, \dots, X_n) \in R$. The **power function** $\beta(\theta)$ is

$$\beta(\theta) = P_{\theta}((X_1, \dots, X_n) \in R)$$

Example

X_1, \dots, X_n iid from a population with mean μ and variance σ^2 .

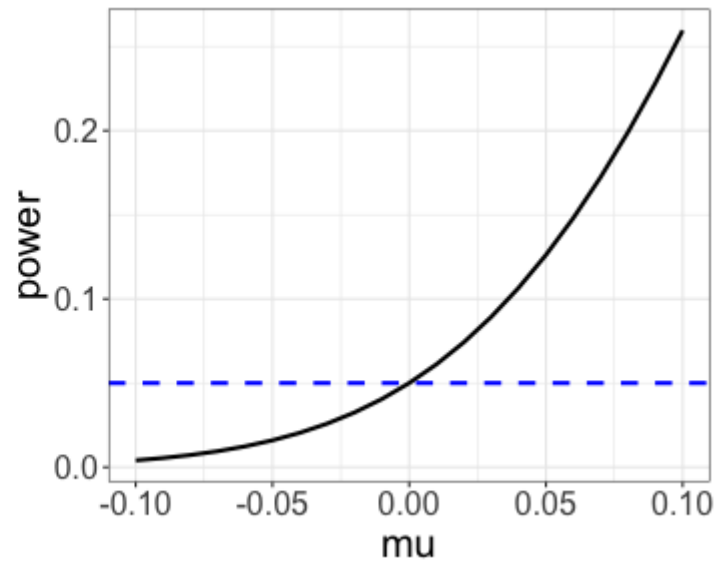
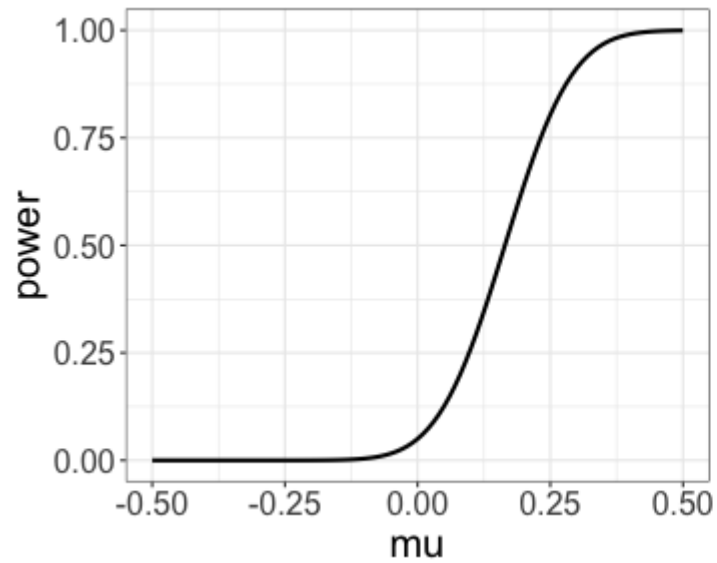
$$H_0 : \mu = \mu_0 \quad H_A : \mu > \mu_0$$

Class activity

$$\beta(\mu) \approx 1 - \Phi \left(z_\alpha - \frac{(\mu - \mu_0)}{\sigma/\sqrt{n}} \right)$$

- + Suppose that $\mu_0 = 0$, $n = 100$, and $\sigma = 1$. Make a plot of $\beta(\mu)$ vs. μ for $\alpha = 0.05$.
- + Now consider testing $H_0 : \mu \leq \mu_0$ vs. $H_A : \mu > \mu_0$. Will this change our rejection region if we want a size α test?

Class activity



Simulation verification

```
n <- 100
sigma <- 1
mu0 <- 0
nreps <- 5000
test_stats <- rep(0, nreps)
for(i in 1:nreps){
  x <- rnorm(n, mu0, sigma)
  test_stats[i] <- (mean(x) - mu0)/(sigma/sqrt(n))
}
mean(test_stats > qnorm(0.05, lower.tail=F))
```

```
## [1] 0.0548
```

Simulation verification

```
n <- 100
sigma <- 1
mu0 <- 0
nreps <- 5000
test_stats <- rep(0, nreps)
for(i in 1:nreps){
  x <- rnorm(n, 0.1, sigma)
  test_stats[i] <- (mean(x) - mu0)/(sigma/sqrt(n))
}
mean(test_stats > qnorm(0.05, lower.tail=F))
```

```
## [1] 0.2646
```