

# Wald tests

# Where we're going

So far:

- ① How can we estimate parameters / fit a model?
  - Maximum likelihood estimation
  - Fisher scoring
- ② Was our fitted model good?
  - Logistic regression diagnostics

Currently:

- ③ How can we use our fitted model for inference?
  - Convergence of MLEs
  - Wald tests

Coming up:

- ④ How else can we test hypotheses? What about confidence intervals?
- ⑤ Why did we focus so much on MLEs?

## Formal definition

Wald test for one parameter

Let  $\theta \in \mathbb{R}$  be a parameter of interest, and let  $\hat{\theta}_n$  be an estimator for which

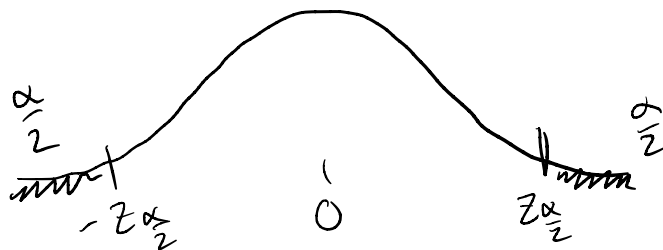
$$\frac{\hat{\theta}_n - \theta}{S_n} \xrightarrow{d} N(0,1) \quad , \text{ for some sequence } S_n$$

$$(S_n^2 \propto \text{Var}(\hat{\theta}_n))$$

To test  $H_0: \theta = \theta_0$  vs.  $H_A: \theta \neq \theta_0$

Let  $Z_n = \frac{\hat{\theta}_n - \theta_0}{S_n}$ . The Wald test rejects  $H_0$

when  $|Z_n| > z_{\frac{\alpha}{2}}$  where  $z_{\frac{\alpha}{2}}$  is the upper  $\frac{\alpha}{2}$  quantile of  $N(0,1)$

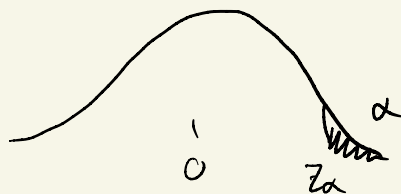


### Comments

- Under  $H_0$  ( $\theta = \theta_0$ ),  $Z_n \approx N(0, 1)$  (if  $n$  is sufficiently large)

$$\begin{aligned} \text{So } P(\text{reject } H_0 \mid H_0 \text{ is true}) &\approx P(|Z| > z_{\frac{\alpha}{2}}) \quad Z \sim N(0, 1) \\ &= \alpha \end{aligned}$$

- To test  $H_A: \theta > \theta_0$ , reject when  $Z_n > z_\alpha$



$$\begin{aligned} &\Rightarrow P(\text{reject } H_0 \mid H_0 \text{ is true}) \\ &= P(Z > z_\alpha) = \alpha \end{aligned}$$

- To test  $H_A: \theta < \theta_0$ , reject when  $Z_n < -z_\alpha$

- Any asymptotically normal statistic can be used to construct a Wald test

# Hypothesis tests for a population mean

Let  $Y_1, Y_2, \dots, Y_n$  be an iid sample from a population with mean  $\mu$  and variance  $\sigma^2$ . We want to test

$$H_0 : \mu = \mu_0 \quad H_A : \mu \neq \mu_0$$

CLT :  $\frac{\bar{Y}_n - \mu}{\sigma/\sqrt{n}} \xrightarrow{d} N(0,1)$

wald test :  $Z_n = \frac{\bar{Y}_n - \mu_0}{\sigma/\sqrt{n}}$

what if  $\sigma$  is unknown?

$$Z_n = \frac{\bar{Y}_n - \mu_0}{\hat{\sigma}/\sqrt{n}} \xrightarrow{d} N(0,1) \quad (\text{under } H_0)$$

$$\hat{\sigma} = \sqrt{\frac{1}{n-1} \sum_{i=1}^n (Y_i - \bar{Y}_n)^2}$$

$$\text{or } \hat{\sigma} = \sqrt{\frac{1}{n} \sum_{i=1}^n (Y_i - \bar{Y}_n)^2}$$

Proof: Slutsky's  
 $\left( \frac{\hat{\sigma}}{\sigma} \xrightarrow{P} 1 \right)$

# Hypothesis tests for a population proportion

Let  $Y_1, Y_2, \dots, Y_n \stackrel{iid}{\sim} \text{Bernoulli}(p)$ . We want to test

$$H_0 : p = p_0 \quad H_A : p \neq p_0$$

What is our Wald test statistic?

$$Z_n = \frac{\hat{p} - p_0}{\sqrt{\frac{\hat{p}(1-\hat{p})}{n}}}$$

$$\hat{p} = \frac{1}{n} \sum_{i=1}^n Y_i = \bar{Y}_n$$

$$\text{var}(Y_i) = p(1-p) \Rightarrow \text{CLT: } \frac{\hat{p} - p}{\sqrt{\frac{p(1-p)}{n}}} \xrightarrow{d} N(0, 1)$$

Alternatively: 
$$\frac{\hat{p} - p_0}{\sqrt{\frac{p_0(1-p_0)}{n}}}$$
 under  $H_0$ , both  $\xrightarrow{d} N(0, 1)$

# Testing multiple parameters

Logistic regression model for the dengue data:

$$Y_i \sim \text{Bernoulli}(p_i)$$

$$\log\left(\frac{p_i}{1-p_i}\right) = \beta_0 + \beta_1 WBC_i + \beta_2 PLT_i$$

Researchers want to know if there is any relationship between white blood cell count or platelet count, and the probability a patient has dengue. What hypotheses should they test?

$$H_0: \beta_1 = \beta_2 = 0$$

$$H_A: \text{at least one of } \beta_1, \beta_2 \neq 0$$

# Testing multiple parameters

```
m1 <- glm(Dengue ~ WBC + PLT, data = dengue,  
          family = binomial)  
summary(m1)
```

```
...  
##              Estimate Std. Error z value Pr(>|z|)  
## (Intercept)  2.6415063  0.1213233   21.77  <2e-16 ***  
## WBC         -0.2892904  0.0134349  -21.53  <2e-16 ***  
## PLT         -0.0065615  0.0005932  -11.06  <2e-16 ***  
## ---  
...
```

Can the researchers test their hypotheses using this output?



## Wald tests for multiple parameters

$$\mathcal{I}^{-1}(\beta) = \begin{bmatrix} \text{Var}(\hat{\beta}_0) & \dots \\ \vdots & \boxed{\text{Var}(\hat{\beta}_1)} & \vdots \\ \vdots & \vdots & \ddots \end{bmatrix}$$

$C \mathcal{I}^{-1}(\beta) C^T$

We know  $\hat{\beta} = \begin{bmatrix} \hat{\beta}_0 \\ \hat{\beta}_1 \\ \hat{\beta}_2 \end{bmatrix} \approx N(\beta, \mathcal{I}^{-1}(\beta))$

We want to test  $\begin{bmatrix} \beta_1 \\ \beta_2 \end{bmatrix}$

Notice that  $\begin{bmatrix} \beta_1 \\ \beta_2 \end{bmatrix} = \underbrace{\begin{bmatrix} 0 & 1 & 0 \\ 0 & 0 & 1 \end{bmatrix}}_C \begin{bmatrix} \beta_0 \\ \beta_1 \\ \beta_2 \end{bmatrix} = C\beta$

$$\Rightarrow \begin{bmatrix} \hat{\beta}_1 \\ \hat{\beta}_2 \end{bmatrix} = C\hat{\beta}$$

$$C\hat{\beta} \approx N(C\beta, C\mathcal{I}^{-1}(\beta)C^T)$$

$$\begin{bmatrix} \hat{\beta}_1 \\ \hat{\beta}_2 \end{bmatrix} \approx \begin{bmatrix} \beta_1 \\ \beta_2 \end{bmatrix}$$

if  $Z \sim N(0,1)$ , then  $Z^2 \sim \chi^2_1$  (HW)

$\Rightarrow Z_n = \frac{\hat{\theta} - \theta_0}{S_n}$ , rejecting when  $|Z_n| > z_{\frac{\alpha}{2}}$  is  
equivalent to rejecting when

$$Z_n^2 > \chi^2_{1,\alpha}$$

where  $\chi^2_{1,\alpha}$  = upper  $\alpha$  quantile of  $\chi^2_1$

if  $Z \in \mathbb{R}^q$  and  $Z \sim N(0, I)$ , then  $Z^T Z \sim \chi^2_q$

# Class activity

[https://sta711-s23.github.io/class\\_activities/ca\\_lecture\\_17.html](https://sta711-s23.github.io/class_activities/ca_lecture_17.html)

+ Wald tests for the dengue data